

Beveik nestacionarių procesų dalinių sumų funkcinės ribinės teoremos

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Structure

1 First order autoregressive processes

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AR(1) process

First order autoregressive $AR(1)$ process is generated according to the scheme

$$y_k = \phi y_{k-1} + \varepsilon_k, \quad k \geq 1 \quad (1)$$

where (ε_k) are innovations and ϕ is an unknown parameter.

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where (ε_k) are innovations and ϕ is an unknown parameter.

- 1 if $|\phi| < 1$, then (1) is stationary process ;
- 2 if $|\phi| > 1$, then (1) is explosive process ;
- 3 if $\phi = 1$, then (1) is nonstationary process.

Estimate of the coefficient ϕ

The least-squares estimate (LSE) of ϕ based on observations y_1, \dots, y_n is

$$\hat{\phi} = \frac{\sum_{k=1}^n y_k y_{k-1}}{\sum_{k=1}^n y_{k-1}^2}.$$

Alternative ways are

- Yule Walker equations (method of moments);
- maximum likelihood estimate.

Some notations

- 1 We denote $\xrightarrow[n \rightarrow \infty]{F}$ the convergence in distribution in the metric space F .
- 2 Classical convergence in distribution of a sequence of random variables is denoted by $\xrightarrow[n \rightarrow \infty]{\mathbb{R}}$.
- 3 Convergence in probability is denoted by $\xrightarrow[n \rightarrow \infty]{P}$.
- 4 $W = (W(t), t \in [0, 1])$ is a standard Brownian motion.
- 5 $U_\gamma(s) = \int_0^s e^{\gamma(s-r)} dW(r)$ is an Ornstein-Uhlenbeck process.
- 6 $\mathfrak{N}(0, \sigma^2)$ denotes normal distribution with mean 0 and variance σ^2 .

Asymptotic behaviour of $\hat{\phi}$

When $|\phi| < 1$ it is well known (see, for example, Mann and Wald (1943) and Anderson (1959)) that the standardized LSE is asymptotically normal :

$$\tau_n := \left(\sum_{k=1}^n y_{k-1}^2 \right)^{1/2} (\hat{\phi} - \phi) \xrightarrow[n \rightarrow \infty]{\mathbb{R}} \mathfrak{N}(0, 1).$$

Remark, that with another standardization

$$\sqrt{n}(\hat{\phi} - \phi) \xrightarrow[n \rightarrow \infty]{\mathbb{R}} \mathfrak{N}(0, 1 - \phi^2).$$

Asymptotic behaviour of $\hat{\phi}$

For $\phi > 1$, Anderson (1959) showed that

$$\tau_n := \left(\sum_{k=1}^n y_{k-1}^2 \right)^{1/2} (\hat{\phi} - \phi) \xrightarrow[n \rightarrow \infty]{\mathbb{R}} \mathfrak{N}(0, 1)$$

is true when the ε_k 's are i.i.d. For general ε_k 's he showed that limiting distribution of τ_n may not exist.

Asymptotic behaviour of ϕ

However when $\phi = 1$, the limit distribution of the properly standardized sequence of the least-squares estimators is non-normal, and it was shown by White (1958) (see also Rao (1978)) that

$$\left(\sum_{k=1}^n y_{k-1}^2 \right)^{1/2} (\hat{\phi} - 1) \xrightarrow[n \rightarrow \infty]{\mathbb{R}} \tau := \frac{\frac{1}{2}(W^2(1) - 1)}{\left(\int_0^1 W^2(t) dt \right)^{1/2}}$$

or

$$n(\hat{\phi} - 1) \xrightarrow[n \rightarrow \infty]{\mathbb{R}} \frac{\int_0^1 W(t) dW(t)}{\int_0^1 W^2(t) dt},$$

where $(W(t), 0 \leq t \leq 1)$ is a standard Brownian motion.

ϕ is "close" to 1

$$\textcircled{1} \mathbb{P}(\tau \leq 0) = \mathbb{P}(W^2(1) \leq 1) = 0.684.$$

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- 2 This indicates that

$$\tau_n := \left(\sum_{k=1}^n y_{k-1}^2 \right)^{1/2} (\hat{\phi} - \phi) \xrightarrow[n \rightarrow \infty]{\mathbb{R}} \mathfrak{N}(0, 1). \quad (2)$$

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- 3 Also

$$\left(\sum_{k=1}^n y_{k-1}^2 \right)^{1/2} (\hat{\phi} - 1) \xrightarrow[n \rightarrow \infty]{\mathbb{R}} \frac{\frac{1}{2}(W^2(1) - 1)}{\left(\int_0^1 W^2(t) dt \right)^{1/2}} \quad (3)$$

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- 4 However, neither (2) nor (3) seems to be intuitive approximations.

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Nearly nonstationary process

Suppose we have first-order autoregressive process $(y_{n,k})$ given by

$$y_{n,k} = \phi_n y_{n,k-1} + \varepsilon_k, \quad k \geq 1, \quad n \geq 1, \quad (4)$$

where

- 1 $\phi_n \rightarrow 1$, as $n \rightarrow \infty$,
- 2 (ε_k) is a sequence of i.i.d.random variables with $\mathbb{E}\varepsilon_k = 0$ and $\mathbb{E}\varepsilon_k^2 = 1$,
- 3 $y_{n,1} \dots, y_{n,n}$ are observations and n is a sample size,
- 4 $y_{n,0}$ random variable.

Parametrisation of ϕ_n

Case 1

$\phi_n = e^{\gamma/n}$ with constant $\gamma < 0$.

This parametrisation was suggested by Phillips (1987 m.).

Case 2

$\phi_n = 1 - \frac{\gamma_n}{n}$, $\gamma_n \rightarrow \infty$ as $n \rightarrow \infty$
and $\gamma_n/n \rightarrow 0$, as $n \rightarrow \infty$

This parametrisation was suggested by Phillips and Giraitis (2006 m.)

Another parametrizations :

- 1 $\phi_n = 1 - \frac{\gamma}{n}$, $\gamma > 0$ (Proposed by Chan & Wei (1987), Cox & Llatas (1991), etc.)
- 2 $\phi_n = 1 - \frac{\gamma_n}{n}$, $\gamma_n \rightarrow \gamma$ (Proposed by Andrews & Guggenberger (2007))
- 3 $\phi_n = 1 - \frac{\gamma_n}{n}$, $\gamma_n \rightarrow 0$ (Proposed by Andrews & Guggenberger (2007))

Some results (Phillips (1987))

If (y_k) is a nearly nonstationary process generated by (4) and $\phi_n = e^{\gamma/n}$, $y_{n,0} = o_P(n^{1/2})$ and (ε_k) are i.i.d. random variables with $\mathbb{E}\varepsilon_0 = 0$ and $\mathbb{E}\varepsilon_0^2 = \sigma^2$, then :

$$\begin{aligned}
 n^{-1/2}y_{[nt]} &\xrightarrow[n \rightarrow \infty]{D[0,1]} \sigma U_\gamma(t), \\
 n^{-3/2} \sum_{j=1}^n y_j &\xrightarrow[n \rightarrow \infty]{\mathbb{R}} \sigma \int_0^1 U_\gamma(r) dr, \\
 n^{-2} \sum_{j=1}^n y_j^2 &\xrightarrow[n \rightarrow \infty]{\mathbb{R}} \sigma^2 \int_0^1 U_\gamma^2(r) dr, \\
 n^{-1} \sum_{j=1}^n y_{j-1} \varepsilon_j &\xrightarrow[n \rightarrow \infty]{\mathbb{R}} \sigma^2 \int_0^1 U_\gamma(r) dW(r).
 \end{aligned}$$

Some results (Phillips and Giraitis (2006))

Suppose (y_k) is a nearly nonstationary process generated by (4) and (ε_k) are i.i.d. random variables with $\mathbb{E}\varepsilon_0 = 0$ and $\mathbb{E}\varepsilon_0^2 = \sigma^2$. Under assumptions $n(1 - \phi_n) \rightarrow \infty$, as $n \rightarrow \infty$ and $\mathbb{E}y_{n,0}^2 = o(n^{1/2})$:

$$\begin{aligned} \frac{(1 - \phi_n^2)^{1/2}}{n^{1/2}} \sum_{j=1}^n \varepsilon_j y_{j-1} &\xrightarrow[n \rightarrow \infty]{\mathbb{R}} \mathfrak{N}(0, \sigma^4), \\ \frac{1 - \phi_n^2}{n} \sum_{j=1}^n y_{j-1}^2 &\xrightarrow[n \rightarrow \infty]{\mathbb{P}} \sigma^2, \\ \frac{(1 - \phi_n)}{n^{1/2}} \sum_{j=1}^n y_j &\xrightarrow[n \rightarrow \infty]{\mathbb{R}} \mathfrak{N}(0, \sigma^2). \end{aligned}$$

Some results (Andrews & Guggenberger (2007))

If (y_k) is a nearly nonstationary process generated by (4) and $\phi_n = 1 - \frac{\gamma_n}{n}$, where $\gamma_n \rightarrow \gamma \in [0, \infty)$, (ε_k) are i.i.d. random variables with $\mathbb{E}\varepsilon_0 = 0$ and $\mathbb{E}\varepsilon_0^2 = \sigma^2$, then :

$$\begin{aligned}
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$$\begin{aligned}
 n^{-1/2} y_{[nt]} &\xrightarrow[n \rightarrow \infty]{D[0,1]} \sigma U_\gamma^*(t), \\
 n^{-3/2} \sum_{j=1}^n y_j &\xrightarrow[n \rightarrow \infty]{\mathbb{R}} \sigma \int_0^1 U_\gamma^*(r) dr, \\
 n^{-2} \sum_{j=1}^n y_j^2 &\xrightarrow[n \rightarrow \infty]{\mathbb{R}} \sigma^2 \int_0^1 (U_\gamma^*)^2(r) dr, \\
 n^{-1} \sum_{j=1}^n y_{j-1} \varepsilon_j &\xrightarrow[n \rightarrow \infty]{\mathbb{R}} \sigma^2 \int_0^1 U_\gamma^*(r) dW(r),
 \end{aligned}$$

where $U_\gamma^*(t) = U_\gamma(t) + (2\gamma)^{-1/2} e^{-\gamma t} Z$ and Z is standard normal random variable.

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If (y_k) is a nearly nonstationary process generated by (4) and $\phi_n = 1 - \frac{\gamma_n}{n}$, where $\gamma_n \rightarrow 0$, (ε_k) are i.i.d. random variables with $\mathbb{E}\varepsilon_0 = 0$ and $\mathbb{E}\varepsilon_0^2 = \sigma^2$ and $y_{n,0} = \sum_{j=0}^{\infty} \phi_n^j \varepsilon_{-j}$ then :

$$(2\gamma_n)^{1/2} n^{-3/2} \sum_{j=1}^n y_{n,j-1} \xrightarrow[n \rightarrow \infty]{\mathbb{R}} \sigma Z,$$

$$2\gamma_n n^{-2} \sum_{j=1}^n y_{n,j-1}^2 \xrightarrow[n \rightarrow \infty]{\mathbb{R}} \sigma^2 Z^2,$$

$$(2\gamma_n)^{1/2} n^{-1} \sum_{j=1}^n y_{j-1} \varepsilon_j \xrightarrow[n \rightarrow \infty]{\mathbb{R}} \sigma^2 Z Z^*,$$

where Z and Z^* are independent standard normal random variables.

The asymptotic behaviour of $\hat{\phi}_n$

Phillips (1987) :

$$n(\hat{\phi}_n - \phi_n) \xrightarrow[n \rightarrow \infty]{\mathbb{R}} \frac{\int_0^1 U_\gamma(r) dW(r)}{\int_0^1 U_\gamma^2(r) dr}$$

Phillips and Giraitis (2006) :

$$\frac{n^{1/2}}{(1 - \phi_n^2)^{1/2}} (\hat{\phi}_n - \phi_n) \xrightarrow[n \rightarrow \infty]{\mathbb{R}} \mathfrak{N}(0, 1)$$

The asymptotic behaviour of ϕ_n

Chan and Wei (1987) :

$$\left(\sum_{k=1}^n y_{k-1}^2 \right)^{1/2} (\hat{\phi}_n - \phi_n) \xrightarrow[n \rightarrow \infty]{\mathbb{R}} \mathcal{L}(\gamma)$$

where

$$\mathcal{L}(\gamma) = \frac{\int_0^1 (1 + (e^{2/\gamma} - 1)t)^{-1} W(t) dW(t)}{\left(\int_0^1 (1 + (e^{2/\gamma} - 1)t)^{-2} W^2(t) dt \right)^{1/2}}.$$

Chan (1988) showed that

$$\mathcal{L}(\gamma) \equiv_{\mathcal{D}} \frac{\int_0^1 U_\gamma(r) dW(r)}{\left(\int_0^1 U_\gamma^2(r) dr \right)^{1/2}}.$$

The asymptotic behaviour of ϕ_n

Andrews & Guggenberger (2007), $\phi_n = 1 - \gamma_n/n$ and $\gamma_n \rightarrow 0$,
 $y_{n,0} = \sum_{j=0}^{\infty} \phi_n^j \varepsilon_{-j}$:

$$(2\gamma_n)^{-1/2} n(\hat{\phi}_n - \phi_n) \xrightarrow[n \rightarrow \infty]{\mathbb{R}} C$$

where C is a Cauchy random variable.

Andrews & Guggenberger (2007), $\phi_n = 1 - \gamma_n/n$ and $\gamma_n \rightarrow \gamma \in (0, \infty]$,
 $y_{n,0} = \sum_{j=0}^{\infty} \phi_n^j \varepsilon_{-j}$:

① $\gamma \in (0, \infty)$

$$n(\hat{\phi}_n - \phi_n) \xrightarrow[n \rightarrow \infty]{\mathbb{R}} \frac{\int_0^1 U_{\gamma}^*(t) dW(r)}{\int_0^1 (U_{\gamma}^*)^2(t) dt}$$

② $\gamma = \infty$

$$(1 - \phi_n^2)^{-1/2} n^{1/2} (\hat{\phi}_n - \phi_n) \xrightarrow[n \rightarrow \infty]{\mathbb{R}} Z.$$

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Process built on the y_k 's

We focus on polygonal line processes built on the y_k 's :

$$S_n^{\text{pl}}(t) := \sum_{k=1}^{\lfloor nt \rfloor} y_{k-1} + (nt - \lfloor nt \rfloor)y_{\lfloor nt \rfloor}, \quad t \in [0, 1], \quad n \geq 1. \quad (5)$$

Remark

The definition of the S_n^{pl} is quite unusual with a general term y_{k-1} where one would expect y_k . However, asymptotic results remain true with y_{k-1} replaced by y_k as well.

Function spaces

The polygonal line process S_n^{pl} can be viewed as a random element either in $C[0, 1]$ or in $H_\alpha^o[0, 1]$. Continuous function space $C[0, 1]$ is endowed with the uniform norm

$$\|f\|_\infty = \sup_{0 \leq t \leq 1} |f(t)| \quad f \in C[0, 1].$$

For $\alpha \in (0, 1)$ the Hölder space

$$H_\alpha^o[0, 1] := \left\{ f \in C[0, 1] : \lim_{\delta \rightarrow 0} \omega_\alpha(f, \delta) = 0 \right\},$$

endowed with the norm $\|f\|_\alpha := |f(0)| + \omega_\alpha(f, 1)$, where

$$\omega_\alpha(f, \delta) := \sup_{\substack{s, t \in [0, 1] \\ 0 < t - s < \delta}} \frac{|f(t) - f(s)|}{|t - s|^\alpha}$$

is a separable Banach space.

Process built on the ε_k 's

The polygonal line process built on i.i.d. random variables (ε_j) is

$$W_n^{\text{pl}}(t) = \sum_{j=1}^{[nt]} \varepsilon_j + (nt - [nt])\varepsilon_{[nt]+1}, \quad t \in [0, 1].$$

By classical Donsker-Prohorov invariance principle

$$n^{-1/2} W_n^{\text{pl}} \xrightarrow[n \rightarrow \infty]{\mathcal{C}[0,1]} W.$$

Invariance principle in Hölder space

By the classical Levy's result on the modulus of continuity of W , $W \in H_\alpha^o[0, 1]$ with probability one for every $0 \leq \alpha < 1/2$.
Račkauskas and Suquet (2004) proved that for $0 < \alpha < 1/2$ the convergence

$$n^{-1/2} \sigma^{-1} W_n^{\text{pl}} \xrightarrow[n \rightarrow \infty]{H_\alpha^o[0,1]} W \quad (6)$$

holds if and only if

$$\lim_{t \rightarrow \infty} t^{1/(1/2-\alpha)} P(|\varepsilon_1| \geq t) = 0. \quad (7)$$

Initial condition

Let us associate to each autoregressive process $(y_{n,k})$ satisfying

$$y_{n,k} = \phi_n y_{n,k-1} + \varepsilon_k, \quad k \geq 1, \quad n \geq 1, \quad (8)$$

the process $(y'_{n,k})$ defined by

$$y'_{n,k} = y_{n,k} - \phi_n^k y_{n,0}.$$

Then $(y'_{n,k})$ satisfies (8) with initialization $y'_{n,0} = 0$.

Initial condition

Proposition 1

Let $S_n^{\text{pl}'}$ be the polygonal line process obtained by substituting in (5) the $y_{n,j}$'s by the $y'_{n,j}$'s. Assume that $c_n S_n^{\text{pl}'}$ converges in distribution in $H_\alpha^o[0, 1]$, where the c_n 's are some positive normalizing constants. Then $c_n S_n^{\text{pl}}$ converges in distribution in $H_\alpha^o[0, 1]$ to the same limit provided that

$$\frac{c_n n^\alpha}{1 - \phi_n} y_{n,0} \xrightarrow[n \rightarrow \infty]{P} 0. \quad (9)$$

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Convergence in $C[0, 1]$ and $H_\alpha^o[0, 1]$ spaces

Theorem 1

In the case 1 where (y_k) is generated by (4) with $\phi_n = e^{\gamma/n}$, $\gamma < 0$, suppose that the sequence of polygonal lines $(n^{-1/2}W_n^{pl})$ converges weakly to the standard Brownian motion W either in $C[0, 1]$ or in $H_\alpha^o[0, 1]$ for some $0 < \alpha < 1/2$. Suppose moreover that $y_{n,0} = o_P(n^{1/2})$ or $y_{n,0} = o_P(n^{1/2-\alpha})$ according to the function space considered. Then $n^{-3/2}S_n^{pl}$ converges weakly, as $n \rightarrow \infty$, in the space under consideration to the integrated Ornstein-Uhlenbeck process J defined by :

$$J(t) := \int_0^t U_\gamma(s) ds, \quad 0 \leq t \leq 1, \quad (10)$$

where $U_\gamma(s) = \int_0^s e^{\gamma(s-r)} dW(r)$.

Convergence with i.i.d. innovations

Taking into account the classical Donsker-Prohorov invariance principle and the functional central limit theorem proved by R&S (2004) we have the following corollary.

Corollary 2

Assume that (y_k) is generated by (4) with $\phi_n = e^{\gamma/n}$, $\gamma < 0$ and that the ε_k 's are i.i.d. and centered. Then the weak convergence of $n^{-3/2}S_n^{\text{pl}}$ to J holds

- *in $C[0, 1]$ provided that $\mathbb{E}\varepsilon_1^2 < \infty$ and $y_{n,0} = o_P(n^{1/2})$;*
- *in $H_\alpha^o[0, 1]$ for $0 < \alpha < 1/2$ under condition*

$$\lim_{t \rightarrow \infty} t^{1/(1/2-\alpha)} P(|\varepsilon_1| \geq t) = 0.$$

and $y_{n,0} = o_P(n^{1/2-\alpha})$.

Essential lemma

Lemma 3

Suppose that the process (y_k) is generated by (4) and $\phi_n = 1 - \gamma_n/n$, where (γ_n) is a sequence of nonnegative numbers such that $\gamma_n \rightarrow \infty$ and $\gamma_n/n \rightarrow 0$ as $n \rightarrow \infty$. Suppose, moreover, that $y_0 = 0$. Let $p \geq 2$. Assume that the innovations (ε_k) satisfy

$$\lim_{t \rightarrow \infty} t^p \mathbb{P}(|\varepsilon_1| > t) = 0 \quad \text{if } p > 2,$$
$$\mathbb{E}\varepsilon_1^2 < \infty \quad \text{if } p = 2.$$

For $p \geq 2$, put $\alpha = 1/2 - 1/p$. Then

$$n^{-1/2} \gamma_n^\alpha \max_{1 \leq k \leq n} |y_k| \xrightarrow[n \rightarrow \infty]{\mathbb{P}} 0.$$

Convergence in $C[0, 1]$

Theorem 4

Suppose that the process (y_k) is generated by (4) and $\phi_n = 1 - \gamma_n/n$, where (γ_n) is a sequence of nonnegative numbers such that $\gamma_n \rightarrow \infty$ and $\gamma_n/n \rightarrow 0$ as $n \rightarrow \infty$. Assume also, that innovations (ε_k) are i.i.d. with $\mathbb{E}\varepsilon_1 = 0$ and $\mathbb{E}\varepsilon_1^2 = 1$ and $y_0 = o_P(n^{1/2})$. Then

$$n^{-1/2}(1 - \phi_n)S_n^{\text{pl}} \xrightarrow[n \rightarrow \infty]{C[0,1]} W.$$

Convergence in H_β^0

Theorem 5

Suppose that the process (y_k) is generated by (4) and $\phi_n = 1 - \gamma_n/n$, where (γ_n) is a sequence of nonnegative numbers such that $\gamma_n \rightarrow \infty$ and $\gamma_n/n \rightarrow 0$ as $n \rightarrow \infty$. Assume also, that innovations (ε_k) are i.i.d. and satisfy condition

$$\lim_{t \rightarrow \infty} t^p \mathbb{P}(|\varepsilon_1| > t) = 0 \quad \text{if } p > 2$$

for some $p > 2$. Put $\alpha = 1/2 - 1/p$. Then for $0 < \beta < \alpha$,

$$n^{-1/2}(1 - \phi_n)S_n^{p1} \xrightarrow[n \rightarrow \infty]{H_\beta^0[0,1]} W$$

provided that $y_0 = o_P(n^{1/2-\beta})$ and

$$\liminf_{n \rightarrow \infty} \gamma_n n^{\beta/\alpha} > 0.$$

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Convergence in $C[0, 1]$ and $H_\alpha^\circ[0, 1]$ spaces

Theorem 1

In the case 1 where (y_k) is generated by (4) with $\phi_n = e^{\gamma/n}$, $\gamma < 0$, suppose that the sequence of polygonal lines $(n^{-1/2}W_n^{pl})$ converges weakly to the standard Brownian motion W either in $C[0, 1]$ or in $H_\alpha^\circ[0, 1]$ for some $0 < \alpha < 1/2$. Suppose moreover that $y_{n,0} = o_P(n^{1/2})$ or $y_{n,0} = o_P(n^{1/2-\alpha})$ according to the function space considered. Then $n^{-3/2}S_n^{pl}$ converges weakly, as $n \rightarrow \infty$, in the space under consideration to the integrated Ornstein-Uhlenbeck process J defined by :

$$J(t) := \int_0^t U_\gamma(s) ds, \quad 0 \leq t \leq 1, \quad (10)$$

where $U_\gamma(s) = \int_0^s e^{\gamma(s-r)} dW(r)$.

Ideas of the proof 1

- 1 Since the Banach spaces $(C[0, 1], \|\cdot\|_\infty)$ and $(H_0^\alpha, \|\cdot\|_0)$ are isomorphic, the unified proof proposed here for the spaces $H_\alpha^\circ[0, 1]$, $0 \leq \alpha < 1/2$, includes the special case of the space $C[0, 1]$.
- 2 Using Proposition 1 and our assumption $y_{n,0} = o_P(n^{1/2-\alpha})$, it is enough to give the proof in the case where $y_{n,0} = 0$.
- 3 Hölder norms of polygonal line is reached at two vertices.

Ideas of the proof 2

Lemma 6

Let $\rho : [0, 1] \rightarrow \mathbb{R}$ be a weight function satisfying the following properties.

- i) ρ is concave.
- ii) $\rho(0) = 0$ and ρ is positive on $(0, 1]$.
- iii) ρ is non decreasing on $[0, 1]$.

Let $t_0 = 0 < t_1 < \dots < t_n = 1$ be a partition of $[0, 1]$ and f be a real valued polygonal line function on $[0, 1]$ with vertices at the t_i 's, i.e. f is continuous on $[0, 1]$ and its restriction to each interval $[t_i, t_{i+1}]$ is an affine function. Define

$$R(s, t) := \frac{|f(t) - f(s)|}{\rho(t - s)}, \quad 0 \leq s < t \leq 1.$$

Then

$$\sup_{0 \leq s < t \leq 1} R(s, t) = \max_{0 \leq i < j \leq n} R(t_i, t_j). \quad (11)$$

Ideas of the proof 3

The idea is to approximate the polygonal line $n^{-3/2}S_n^{\text{pl}}$ by some linear interpolation of a smooth process J_n

$$J_n(t) := \int_0^t n^{-1/2} W_n^{\text{pl}}(s) ds + \gamma \int_0^t \int_0^s e^{\gamma(s-r)} n^{-1/2} W_n^{\text{pl}}(r) dr ds.$$

which is a functional of $n^{-1/2} W_n^{\text{pl}}$, continuous in Hölder topology, with $\|n^{-3/2} S_n^{\text{pl}} - J_n\|_\alpha = o_P(1)$.

As the functional

$$H_\alpha^o[0, 1] \rightarrow H_\alpha^o[0, 1] \quad : \quad x \mapsto \int_0^\bullet x(s) ds + \gamma \int_0^\bullet \int_0^s e^{\gamma(s-r)} x(r) dr ds$$

is continuous on $H_\alpha^o[0, 1]$, the convergence of J_n to J follows from the convergence of $n^{-1/2} W_n^{\text{pl}}$ to W .

Convergence in $C[0, 1]$

Theorem 4

Suppose that the process (y_k) is generated by (4) and $\phi_n = 1 - \gamma_n/n$, where (γ_n) is a sequence of nonnegative numbers such that $\gamma_n \rightarrow \infty$ and $\gamma_n/n \rightarrow 0$ as $n \rightarrow \infty$. Assume also, that innovations (ε_k) are i.i.d. with $\mathbb{E}\varepsilon_1 = 0$ and $\mathbb{E}\varepsilon_1^2 = 1$ and $y_0 = o_P(n^{1/2})$. Then

$$n^{-1/2}(1 - \phi_n)S_n^{\text{pl}} \xrightarrow[n \rightarrow \infty]{C[0,1]} W.$$

Ideas of the proof

- Using Proposition 1 and assumption $y_0 = o_P(n^{1/2})$, it suffices to prove the result when $y_0 = 0$;
- In view of the Donsker-Prohorov invariance principle, it suffices to show that

$$\left\| n^{-1/2}(1 - \phi_n)S_n^{\text{pl}} - n^{-1/2}W_n^{\text{pl}} \right\|_{\infty} \xrightarrow[n \rightarrow \infty]{P} 0.$$

- As supremum norm is reached at one of its vertices and

$$(1 - \phi_n) \sum_{j=1}^k y_{j-1} = -y_k + \sum_{j=1}^k \varepsilon_j$$

so it is enough to show that

$$n^{-1/2} \max_{1 \leq k \leq n} |y_k| \xrightarrow[n \rightarrow \infty]{P} 0. \quad (12)$$

- And (12) is true because of Lemma 1 with $p = 2$.

Convergence in H_β^0

Theorem 5

Suppose that the process (y_k) is generated by (4) and $\phi_n = 1 - \gamma_n/n$, where (γ_n) is a sequence of nonnegative numbers such that $\gamma_n \rightarrow \infty$ and $\gamma_n/n \rightarrow 0$ as $n \rightarrow \infty$. Assume also, that innovations (ε_k) are i.i.d. and satisfy condition

$$\lim_{t \rightarrow \infty} t^p \mathbb{P}(|\varepsilon_1| > t) = 0 \quad \text{if } p > 2$$

for some $p > 2$. Put $\alpha = 1/2 - 1/p$. Then for $0 < \beta < \alpha$,

$$n^{-1/2}(1 - \phi_n)S_n^{p1} \xrightarrow[n \rightarrow \infty]{H_\beta^0[0,1]} W$$

provided that $y_0 = o_P(n^{1/2-\beta})$ and

$$\liminf_{n \rightarrow \infty} \gamma_n n^{\beta/\alpha} > 0.$$

Ideas of the proof

- 1 In view of Hölderian invariance principle it is enough to show that

$$\left\| n^{-1/2}(1 - \phi_n)S_n^{\text{Pl}} - n^{-1/2}W_n^{\text{Pl}} \right\|_{\beta} \xrightarrow[n \rightarrow \infty]{\text{P}} 0.$$

- 2 As Hölderian norms of the polygonal line process is reached at two vertices, so the proof reduces to

$$n^{\beta-1/2} \max_{1 \leq k \leq n} |y_k| \xrightarrow[n \rightarrow \infty]{\text{P}} 0. \quad (13)$$

- 3 The convergence (13) holds due to Proposition 1 and Lemma 1 with condition

$$\limsup_{n \rightarrow \infty} n^{\beta} / \gamma_n^{\alpha} < \infty.$$

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