

Functional central limit theorems for sums of nearly nonstationary processes

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11th of April, 2012

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AR(1) process

First order autoregressive $AR(1)$ process is generated according to the scheme

$$y_k = \phi y_{k-1} + \varepsilon_k, \quad k = 1, 2, \dots, n, \quad y_0 = 0. \quad (1)$$

where $(\varepsilon_k)_{k \geq 0}$ are the innovations at time k , and ϕ is an unknown parameter.

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- 1 $|\phi| < 1$ - (1) is stationary process ;
- 2 $|\phi| > 1$ - (1) is explosive process ;
- 3 $\phi = 1$ - (1) is nonstationary process.

Asymptotic behaviour of ϕ

The least-squares estimator (LSE) of ϕ is

$$\hat{\phi} = \frac{\sum_{k=1}^n y_k y_{k-1}}{\sum_{k=1}^n y_{k-1}^2}.$$

When $|\phi| < 1$ it is well known (see, for example, Mann and Wald (1943) and Anderson (1959)) that the standardized LSE is asymptotically normal :

$$\tau_n := \left(\sum_{k=1}^n y_{k-1}^2 \right)^{1/2} (\hat{\phi} - \phi) \xrightarrow[n \rightarrow \infty]{\mathbb{R}} \mathfrak{N}(0, 1).$$

Here $\mathfrak{N}(0, \sigma^2)$ denotes normal distribution with mean 0 and variance σ^2 .

Asymptotic behaviour of ϕ

For $\phi > 1$, Anderson (1959) showed that

$$\tau_n := \left(\sum_{k=1}^n y_{k-1}^2 \right)^{1/2} (\hat{\phi} - \phi) \xrightarrow[n \rightarrow \infty]{\mathbb{R}} \mathfrak{N}(0, 1)$$

is true when the ε_k 's are i.i.d. For general ε_k 's he showed that limiting distribution of τ_n may not exist.

The classical convergence in distribution of a sequence of random variables is denoted by $\xrightarrow[n \rightarrow \infty]{\mathbb{R}}$.

Asymptotic behaviour of ϕ

However when $\phi = 1$, the limit distribution of the properly standardized sequence of the least-squares estimators is non-normal, and it was shown by White (1958) (see also Rao (1978)) that

$$\left(\sum_{k=1}^n y_{k-1}^2 \right)^{1/2} (\hat{\phi} - 1) \xrightarrow[n \rightarrow \infty]{\mathbb{R}} \tau := \frac{\frac{1}{2}(W^2(1) - 1)}{\left(\int_0^1 W^2(t) dt \right)^{1/2}}$$

where $(W(t), 0 \leq t \leq 1)$ is a standard Brownian motion.

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- 3 Also

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could be used to approximate the distribution of τ_n when ϕ is close to one (Evans and Savin (1981)).

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- 4 However, neither (2) nor (3) seems to be intuitive approximations.

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Nearly nonstationary process

Suppose we have first-order autoregressive process $(y_{n,k} | k = 1, \dots, n; n = 1, 2, \dots)$ given by

$$y_{n,k} = \phi_n y_{n,k-1} + \varepsilon_k, \quad y_{n,0} = y_0 = 0 \quad (4)$$

where

- 1 $\phi_n \rightarrow 1$, as $n \rightarrow \infty$,
- 2 $(\varepsilon_k)_{k \geq 0}$ is a sequence of i.i.d. random variables with $\mathbb{E}\varepsilon_k = 0$ and $\mathbb{E}\varepsilon_k^2 = 1$,
- 3 n is a sample size.

Parametrisation of ϕ_n

Case 1

$\phi_n = e^{\gamma/n}$ with constant $\gamma < 0$.
This parametrisation was suggested
by Phillips (1987 m.).

Case 2

$\phi_n = 1 - \frac{\gamma_n}{n}$, $\gamma_n \rightarrow \infty$ slower than n .
This parametrisation was suggested
by Phillips and Giraitis (2006 m.)

Some results (Phillips (1987 m.))

If $(y_k)_{k \geq 0}$ is a nearly non stationary process generated by (4) and $\phi_n = e^{\gamma/n}$, $(\varepsilon_k)_{k \geq 0}$ are i.i.d. random variables with $\mathbb{E}\varepsilon_0 = 0$ and $\mathbb{E}\varepsilon_0^2 = \sigma^2$, then :

$$\begin{aligned}
 n^{-1/2} y_{[nt]} &\xrightarrow[n \rightarrow \infty]{D[0,1]} \sigma U_\gamma(t), \\
 n^{-3/2} \sum_{j=1}^n y_j &\xrightarrow[n \rightarrow \infty]{\mathbb{R}} \sigma \int_0^1 U_\gamma(r) dr, \\
 n^{-2} \sum_{j=1}^n y_j^2 &\xrightarrow[n \rightarrow \infty]{\mathbb{R}} \sigma^2 \int_0^1 U_\gamma^2(r) dr, \\
 n^{-1} \sum_{j=1}^n y_{j-1} \varepsilon_j &\xrightarrow[n \rightarrow \infty]{\mathbb{R}} \sigma^2 \int_0^1 U_\gamma(r) dW(r).
 \end{aligned}$$

where

$$U_\gamma(t) = \int_0^t e^{(t-s)\gamma} dW(s) \quad 0 \leq t \leq 1$$

Some results (Phillips and Giraitis (2006 m.))

Suppose $(y_k)_{k \geq 0}$ is a nearly non stationary process generated by (4) and $(\varepsilon_k)_{k \geq 0}$ are i.i.d. random variables with $\mathbb{E}\varepsilon_0 = 0$ and $\mathbb{E}\varepsilon_0^2 = \sigma^2$. Under assumption : $n(1 - \phi_n) \rightarrow \infty$, as $n \rightarrow \infty$:

$$\frac{(1 - \phi_n^2)^{1/2}}{n^{1/2}} \sum_{j=1}^n \varepsilon_j y_{j-1} \xrightarrow[n \rightarrow \infty]{\mathbb{R}} \mathfrak{N}(0, \sigma^4),$$

$$\frac{1 - \phi_n^2}{n} \sum_{j=1}^n y_{j-1}^2 \xrightarrow[n \rightarrow \infty]{\mathbb{P}} \sigma^2,$$

$$\frac{(1 - \phi_n)}{n^{1/2}} \sum_{j=1}^n y_j \xrightarrow[n \rightarrow \infty]{\mathbb{R}} \mathfrak{N}(0, \sigma^2).$$

Convergence in probability is denoted by $\xrightarrow[n \rightarrow \infty]{\mathbb{P}}$.

The asymptotic behaviour of $\hat{\phi}_n$

Phillips (1987 m.) :

$$n(\hat{\phi}_n - \phi_n) \xrightarrow[n \rightarrow \infty]{\mathbb{R}} \frac{\int_0^1 U_\gamma(r) dW(r)}{\int_0^1 U_\gamma^2(r) dr}$$

Phillips and Giraitis (2006 m.) :

$$\frac{n^{1/2}}{(1 - \hat{\phi}_n^2)^{1/2}} (\hat{\phi}_n - \phi_n) \xrightarrow[n \rightarrow \infty]{\mathbb{R}} \mathfrak{N}(0, 1)$$
$$\frac{n^{1/2}}{|1 - \hat{\phi}_n^2|^{1/2}} (\hat{\phi}_n - \phi_n) \xrightarrow[n \rightarrow \infty]{\mathbb{R}} \mathfrak{N}(0, 1)$$

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Process built on the y_k 's

We focus on polygonal line processes built on the y_k 's :

$$S_n^{\text{pl}}(t) := \sum_{k=1}^{[nt]} y_{k-1} + (nt - [nt])y_{[nt]}, \quad t \in [0, 1], \quad n \geq 1.$$

Remark

The definition of the S_n^{pl} is quite unusual with a general term y_{k-1} where one would expect y_k . However, asymptotic results remain true with y_{k-1} replaced by y_k as well.

Function spaces

The polygonal line process S_n^{pl} can be viewed as a random element either in $C[0, 1]$ or in $H_\alpha^o[0, 1]$. Continuous function space $C[0, 1]$ is endowed with the uniform norm

$$\|f\|_\infty = \sup_{0 \leq t \leq 1} |f(t)| \quad f \in C[0, 1].$$

For $\alpha \in (0, 1)$ the Hölder space

$$H_\alpha^o[0, 1] := \left\{ f \in C[0, 1] : \lim_{\delta \rightarrow 0} \omega_\alpha(f, \delta) = 0 \right\},$$

endowed with the norm $\|f\|_\alpha := |f(0)| + \omega(f, 1)$, where

$$\omega_\alpha(f, \delta) := \sup_{\substack{s, t \in [0, 1] \\ 0 < t - s < \delta}} \frac{|f(t) - f(s)|}{|t - s|^\alpha}$$

is a separable Banach space.

Process built on the ε_k 's

The polygonal line process built on i.i.d. random variables (ε_j) is

$$W_n^{\text{pl}}(t) = \sum_{j=1}^{[nt]} \varepsilon_j + (nr - [nt])\varepsilon_{[nt]+1}, \quad t \in [0, 1].$$

By classical Donsker-Prohorov invariance principle

$$n^{-1/2} W_n^{\text{pl}} \xrightarrow[n \rightarrow \infty]{\mathcal{C}[0,1]} W.$$

In what follows $\xrightarrow[n \rightarrow \infty]{\mathbb{E}}$ signifies convergence in distribution in the metric space \mathbb{E} .

Invariance principle in Hölder space

By the classical Levy's result on the modulus of continuity of W , $W \in H_\alpha^0[0, 1]$ with probability one for every $0 \leq \alpha < 1/2$.
Račkauskas and Suquet (2004) proved that for $0 < \alpha < 1/2$ the convergence

$$n^{-1/2} \sigma^{-1} W_n^{\text{pl}} \xrightarrow[n \rightarrow \infty]{H_\alpha^0[0,1]} W$$

holds if and only if

$$\lim_{t \rightarrow \infty} t^{1/(1/2-\alpha)} P(|\varepsilon_1| \geq t) = 0. \quad (5)$$

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Convergence in $C[0, 1]$ and $H_\alpha^\circ[0, 1]$ spaces

Theorem

Suppose that (y_k) is generated by (4), $\phi_n = e^{\gamma/n}$ with $\gamma < 0$ and that the sequence of polygonal lines $n^{-1/2}W_n^{\text{pl}}$ converges weakly to the standard Brownian motion W either in $C[0, 1]$ or in $H_\alpha^\circ[0, 1]$ for some $0 < \alpha < 1/2$. Then $n^{-3/2}S_n^{\text{pl}}$ converges weakly in the space under consideration to the integrated Ornstein-Uhlenbeck process J defined by :

$$J(t) := \int_0^t U_\gamma(s) ds, \quad 0 \leq t \leq 1, \quad (6)$$

where $U_\gamma(s) = \int_0^s e^{\gamma(s-r)} dW(r)$.

Convergence with i.i.d. innovations

Taking into account the classical Donsker-Prohorov invariance principle and the functional central limit theorem proved by R&S (2004) we have the following corollary.

Corollary

When the ε_k 's are i.i.d. and centered, the weak convergence of $n^{-3/2}S_n^{p1}$ to J holds

- *in $C[0, 1]$ provided that $\mathbb{E}\varepsilon_1^2 < \infty$;*
- *in $H_\alpha^0[0, 1]$ for $0 < \alpha < 1/2$ under condition (5).*

$$y_{n,0} \neq 0$$

Remark

For the sake of simplicity, we assumed the initialization $y_{n,0} = y_0 = 0$. Nevertheless if $y_{n,0} \neq 0$, then

$$y_k = \sum_{j=1}^k e^{(k-j)\gamma/n} \varepsilon_j + e^{k\gamma/n} y_{n,0}$$

and because of the normalization $n^{-3/2}$ the proof of the theorem holds, provided that $y_{n,0} = o_P(n^{3/2})$.

Convergence $C[0, 1]$ space

Theorem

Suppose $(y_k)_{k \geq 0}$ is generated by (4) and $\phi_n = 1 - \gamma_n/n$, where (γ_n) is a sequence of non negative numbers and tends to infinity slower than n . Assume also that the innovations $(\varepsilon_k)_{k \geq 0}$ are i.i.d. with $\mathbb{E}\varepsilon_k = 0$, $\mathbb{E}\varepsilon_k^2 = 1$. Then the following convergence holds.

$$n^{-1/2}(1 - \phi_n)S_n^{\text{pl}} \xrightarrow[n \rightarrow \infty]{C[0,1]} W.$$

Convergence in $H_\alpha^0[0, 1]$ space

Theorem

Suppose $(y_k)_{k \geq 0}$ is generated by (4) and $\phi_n = 1 - \gamma_n/n$, where (γ_n) is a sequence of non negative numbers and tends to infinity slower than n . Assume also that the innovations $(\varepsilon_k)_{k \geq 0}$ are i.i.d. and satisfy condition $\lim_{t \rightarrow \infty} t^p P(|\varepsilon_0| > t) = 0$ for some $p > 2$. Put $\alpha = \frac{1}{2} - \frac{1}{p}$. Then for $0 < \beta < \alpha$,

$$n^{-1/2}(1 - \phi_n)S_n^{\text{pl}} \xrightarrow[n \rightarrow \infty]{H_\beta^0[0,1]} W,$$

provided that

$$\liminf_{n \rightarrow \infty} \gamma_n n^{-\frac{\beta}{\alpha}} > 0.$$

Essential lemma

A key point in all the following limit theorems is to keep a good control on the asymptotic behaviour of $\max_{1 \leq k \leq n} |y_k|$.

Lemma

Let $p \geq 2$. Assume that the innovations $(\varepsilon_k)_{k \geq 0}$ satisfy

$$\begin{aligned} \lim_{t \rightarrow \infty} t^p P(|\varepsilon_0| > t) &= 0, & \text{if } p > 2 \\ \mathbb{E} \varepsilon_0^2 < \infty & & \text{if } p = 2 \end{aligned}$$

For $p \geq 2$, put $\alpha = 1/2 - 1/p$. Then

$$n^{-1/2} \gamma_n^\alpha \max_{1 \leq k \leq n} |y_k| \xrightarrow[n \rightarrow \infty]{P} 0.$$

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Residuals process

In further research we focus on processes built from $\widehat{\varepsilon}_k$:

$$\widehat{Z}_n^{\text{pl}}(t) = \sum_{j=1}^{[nt]} \widehat{\varepsilon}_j + (nr - [nt])\widehat{\varepsilon}_{[nt]+1} \quad (7)$$

where $\widehat{\varepsilon}_k$ are residuals of the process y_k defined by

$$\widehat{\varepsilon}_k = y_k - \widehat{\phi}_n y_{k-1}$$

Convergence in case 1

Theorem

Let $\alpha \in (0, 1/2)$. Suppose that y_k is generated by (4), $\phi_n = e^{\gamma/n}$. Also (ε_k) are independent identically distributed random variables with $\mathbb{E}\varepsilon_0 = 0$ and $\mathbb{E}\varepsilon_0^2 = 1$. Then

$$n^{-1/2} \widehat{Z}_n^{\text{pl}} \xrightarrow[n \rightarrow \infty]{H_\alpha^0[0,1]} W - A^{-1}BJ, \quad (8)$$

if and only if condition (5) holds.

Here $B = \int_0^1 U_\gamma(r) dW(r)$, $A = \int_0^1 U_\gamma(r)^2 dr$ and $J(t) := \int_0^t U_\gamma(s) ds$.

Convergence in case 2 in $C[0, 1]$

Theorem

Suppose y_k is generated by (4) and $\phi_n = 1 - \gamma_n/n$, where γ_n is non negative and goes to infinity slower than n . Assume also that the innovations (ε_k) are i.i.d. with $\mathbb{E}\varepsilon_k = 0$, $\mathbb{E}\varepsilon_k^2 = 1$. Then the following convergences hold :

$$n^{-1/2} \widehat{Z}_n^{\text{pl}} \xrightarrow[n \rightarrow \infty]{C[0,1]} W. \quad (9)$$

Convergence in case 2 in $H_\alpha^0[0, 1]$

Theorem

Suppose y_k is generated by (4) and $\phi_n = 1 - \gamma_n/n$, where γ_n is non negative and goes to infinity slower than n . Assume also that the innovations $(\varepsilon_k)_{k \geq 1}$ are i.i.d. and satisfy condition

$$\lim_{t \rightarrow \infty} t^p \mathbb{P}(|\varepsilon_0| > t) = 0 \quad (10)$$

for some $p > 2$. Put $\alpha = \frac{1}{2} - \frac{1}{p}$. Then for $0 < \beta \leq \alpha$,

$$n^{-1/2} \widehat{Z}_n^{p1} \xrightarrow[n \rightarrow \infty]{H_\beta^0[0,1]} W, \quad (11)$$

where W is a standard Wiener process if

$$\liminf_{n \rightarrow \infty} \gamma_n n^{-\frac{2\beta}{1+2\alpha}} > 0. \quad (12)$$

Possible applications

- 1 Epidemic change in mean of innovations ;
- 2 Epidemic change in variance of innovations ;
- 3 Epidemic change in parameter γ ;
- 4 Epidemic change in coefficient ϕ_n between stationary and nonstationary state.

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