

Functional central limit theorems for nearly nonstationary processes and applications for testing epidemic change

Jurgita Markevičiūtė^{a,b}
Alfredas Račkauskas^a, Charles Suquet^b

^aDepartment of Mathematics and Informatics, Vilnius University,

^bLaboratoire P. Painlevé, UMR 8524 CNRS Université Lille I

11th of June, 2012

Table of contents

- 1 First order autoregressive processes
- 2 Notations and Background
- 3 Function central limit theorems
- 4 Applications for testing epidemic change

AR(1) process

First order autoregressive $AR(1)$ process is generated according to the scheme

$$y_k = \phi y_{k-1} + \varepsilon_k, \quad k \geq 1, \quad (1)$$

where (ε_k) are the innovations at time k , and ϕ is an unknown parameter.

AR(1) process

First order autoregressive $AR(1)$ process is generated according to the scheme

$$y_k = \phi y_{k-1} + \varepsilon_k, \quad k \geq 1, \quad (1)$$

where (ε_k) are the innovations at time k , and ϕ is an unknown parameter.

- 1 $|\phi| < 1$, then (1) is stationary process

AR(1) process

First order autoregressive $AR(1)$ process is generated according to the scheme

$$y_k = \phi y_{k-1} + \varepsilon_k, \quad k \geq 1, \quad (1)$$

where (ε_k) are the innovations at time k , and ϕ is an unknown parameter.

- 1 $|\phi| < 1$, then (1) is stationary process ;
- 2 $|\phi| > 1$, then (1) is explosive process

AR(1) process

First order autoregressive $AR(1)$ process is generated according to the scheme

$$y_k = \phi y_{k-1} + \varepsilon_k, \quad k \geq 1, \quad (1)$$

where (ε_k) are the innovations at time k , and ϕ is an unknown parameter.

- 1 $|\phi| < 1$, then (1) is stationary process ;
- 2 $|\phi| > 1$, then (1) is explosive process ;
- 3 $\phi = 1$, then (1) is nonstationary process.

ϕ is "close" to 1

- 1 For $|\phi| \leq 1$ the asymptotics :

$$\left(\sum_{k=1}^n y_{k-1}^2 \right)^{1/2} (\hat{\phi} - \phi) \xrightarrow[n \rightarrow \infty]{\mathbb{R}} \mathfrak{N}(0, 1). \quad (2)$$

- 2 For $\phi = 1$

$$\left(\sum_{k=1}^n y_{k-1}^2 \right)^{1/2} (\hat{\phi} - 1) \xrightarrow[n \rightarrow \infty]{\mathbb{R}} \frac{\frac{1}{2}(W^2(1) - 1)}{\left(\int_0^1 W^2(t) dt \right)^{1/2}} \quad (3)$$

- 3 Because of $\mathbb{P}(\tau \leq 0) = \mathbb{P}(W^2(1) \leq 1) = 0.684$, (2) may not be a satisfactory approximation when ϕ is "close" to 1 and the sample size is moderate. Also (3) could be used. However, neither (2) nor (3) seems to be intuitive approximations.

Nearly nonstationary process

Suppose we have first-order autoregressive process $(y_{n,k})$ given by

$$y_{n,k} = \phi_n y_{n,k-1} + \varepsilon_k, \quad k \geq 1, \quad n \geq 1, \quad (4)$$

where

- 1 $\phi_n \rightarrow 1$, as $n \rightarrow \infty$,
- 2 (ε_k) is a sequence of i.i.d.random variables with $\mathbb{E}\varepsilon_k = 0$ and $\mathbb{E}\varepsilon_k^2 = 1$,
- 3 $y_{n,1} \dots, y_{n,n}$ are observations and n is a sample size,
- 4 for simplicity $y_{n,0} = y_0 = 0$.

Parametrisation of ϕ_n

Case 1

$\phi_n = e^{\gamma/n}$ with constant $\gamma < 0$.
This parametrisation was suggested
by Phillips (1987 m.).

Case 2

$\phi_n = 1 - \frac{\gamma_n}{n}$, $\gamma_n \rightarrow \infty$ slower than n .
This parametrisation was suggested
by Phillips and Giraitis (2006 m.)

We will use LSE estimate

$$\hat{\phi}_n = \frac{\sum_{k=1}^n y_k y_{k-1}}{\sum_{k=1}^n y_{k-1}^2}$$

based on observations y_1, \dots, y_n with $y_0 = 0$ for convenience.

Table of contents

- 1 First order autoregressive processes
- 2 Notations and Background
- 3 Function central limit theorems
- 4 Applications for testing epidemic change

Process built on the y_k 's

We focus on polygonal line processes built on the y_k 's :

$$S_n^{\text{Pl}}(t) := \sum_{k=1}^{[nt]} y_{k-1} + (nt - [nt])y_{[nt]}, \quad t \in [0, 1], \quad n \geq 1.$$

and on polygonal line process built on the $\hat{\varepsilon}_k$'s :

$$\widehat{W}_n^{\text{Pl}}(t) := \sum_{k=1}^{[nt]} \hat{\varepsilon}_k + (nt - [nt])\hat{\varepsilon}_{[nt]+1}, \quad t \in [0, 1], \quad n \geq 1,$$

where $\hat{\varepsilon}_k$ are residuals of the process y_k defined by

$$\hat{\varepsilon}_k = y_k - \hat{\phi}_n y_{k-1}$$

Function spaces

The polygonal line process S_n^{pl} can be viewed as a random element in Hölder space $H_\alpha^o[0, 1]$. For $\alpha \in (0, 1)$

$$H_\alpha^o[0, 1] := \left\{ f \in C[0, 1] : \lim_{\delta \rightarrow 0} \omega_\alpha(f, \delta) = 0 \right\},$$

endowed with the norm $\|f\|_\alpha := |f(0)| + \omega(f, 1)$, where

$$\omega_\alpha(f, \delta) := \sup_{\substack{s, t \in [0, 1] \\ 0 < t - s < \delta}} \frac{|f(t) - f(s)|}{|t - s|^\alpha}$$

is a separable Banach space.

Table of contents

- 1 First order autoregressive processes
- 2 Notations and Background
- 3 **Function central limit theorems**
- 4 Applications for testing epidemic change

Convergence in $H_\alpha^o[0, 1]$ spaces

Theorem

Suppose that (y_k) is generated by (4), $\phi_n = e^{\gamma/n}$ with $\gamma < 0$ and that the sequence of polygonal lines $n^{-1/2} W_n^{\text{pl}}$ converges weakly to the standard Brownian motion W in $H_\alpha^o[0, 1]$ for some $0 < \alpha < 1/2$. Then $n^{-3/2} S_n^{\text{pl}}$ converges weakly in the space under consideration to the integrated Ornstein-Uhlenbeck process J defined by :

$$J(t) := \int_0^t U_\gamma(s) ds, \quad 0 \leq t \leq 1, \quad (5)$$

where $U_\gamma(s) = \int_0^s e^{\gamma(s-r)} dW(r)$.

Convergence in $H_{\beta}^0[0, 1]$ space

Theorem

Suppose (y_k) is generated by (4) and $\phi_n = 1 - \gamma_n/n$, where (γ_n) is a sequence of non negative numbers, $\gamma_n \rightarrow \infty$ and $\gamma_n/n \rightarrow 0$ as $n \rightarrow \infty$. Assume also that the innovations (ε_k) are i.i.d. and satisfy condition $\lim_{t \rightarrow \infty} t^p P(|\varepsilon_0| > t) = 0$ for some $p > 2$. Put $\alpha = \frac{1}{2} - \frac{1}{p}$. Then for $0 < \beta < \alpha$,

$$n^{-1/2}(1 - \phi_n)S_n^{\text{pl}} \xrightarrow[n \rightarrow \infty]{H_{\beta}^0[0,1]} W,$$

provided that

$$\liminf_{n \rightarrow \infty} \gamma_n n^{-\frac{\beta}{\alpha}} > 0.$$

Convergence in $H_\alpha^o[0, 1]$

Theorem

Let $\alpha \in (0, 1/2)$. Suppose that (y_k) is generated by (4), $\phi_n = e^{\gamma/n}$. Also (ε_k) are independent identically distributed random variables with $\mathbb{E}\varepsilon_0 = 0$. Then

$$n^{-1/2} \widehat{W}_n^{\text{pl}} \xrightarrow[n \rightarrow \infty]{H_\alpha^o[0,1]} W - A^{-1} B' J, \quad (6)$$

if and only if condition

$$\lim_{t \rightarrow \infty} t^{1/(1/2-\alpha)} P(|\varepsilon_1| \geq t) = 0. \quad (7)$$

holds.

Here $B' = \int_0^1 U_\gamma(r) dW(r)$, $A = \int_0^1 U_\gamma(r)^2 dr$ and $J(t) := \int_0^t U_\gamma(s) ds$.

Convergence in $H_\alpha^0[0, 1]$

Theorem

Suppose (y_k) is generated by (4) and $\phi_n = 1 - \gamma_n/n$, where γ_n is a sequence of non negative constants, $\gamma_n \rightarrow \infty$ and $\gamma_n/n \rightarrow 0$ as $n \rightarrow \infty$. Assume also that the innovations (ε_k) are i.i.d. and satisfy condition

$$\lim_{t \rightarrow \infty} t^p \mathbb{P}(|\varepsilon_0| > t) = 0 \quad (8)$$

for some $p > 2$. Put $\alpha = \frac{1}{2} - \frac{1}{p}$. Then for $0 < \beta \leq \alpha$,

$$n^{-1/2} \widehat{W}_n^{p1} \xrightarrow[n \rightarrow \infty]{H_\beta^0[0,1]} W, \quad (9)$$

where W is a standard Wiener process if

$$\liminf_{n \rightarrow \infty} \gamma_n n^{-\frac{2\beta}{1+2\alpha}} > 0. \quad (10)$$

Table of contents

- 1 First order autoregressive processes
- 2 Notations and Background
- 3 Function central limit theorems
- 4 Applications for testing epidemic change

Epidemic change in mean

Hypothesis for y_k 's

$$z_{n,k} = a_n \mathbf{1}_{\{k^* < k \leq k^* + l^*\}} + y_{n,k}$$

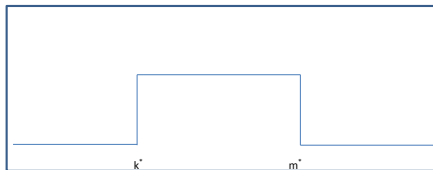
$$H_0 : a_n = 0;$$

$$H_A : a_n \neq 0.$$

Hypothesis for ε_k 's

$$H_0 : \mathbb{E}\varepsilon_k = 0;$$

$$H_A : \mathbb{E}\varepsilon_k = a \mathbf{1}_{\{k^* < k \leq k^* + l^*\}}.$$



Test statistics with y_k 's

We construct uniform increment statistics :

$$U_{IS}(n, \alpha) := \sup_{k^*, l^*} \frac{\left| S_n^{\text{st}}(\mathbb{I}_n^*) - \frac{l^*}{n} S_n^{\text{st}}(n) \right|}{\left| \frac{l^*}{n} \left(1 - \frac{l^*}{n} \right) \right|^\alpha}$$

where

$$\begin{aligned} \mathbb{I}_n^* &= \{k^* + 1, \dots, k^* + l^*\} \\ S_n^{\text{st}}(\mathbb{I}_n^*) &= \sum_{k \in \mathbb{I}_n^*} y_{n, k-1} \\ S_n^{\text{st}}(n) &= \sum_{k=1}^n y_{n, k-1} \end{aligned}$$

Behaviour under null hypothesis in case 1

Theorem

Suppose that (y_k) is generated by (4), $\phi_n = e^{\gamma/n}$ with $\gamma < 0$ and that the sequence (ε_k) are i.i.d. random variables with mean 0 and satisfy condition $\lim_{t \rightarrow \infty} t^{1/(1/2-\alpha)} P(|\varepsilon_1| \geq t) = 0$. Then under hypothesis H_0

$$n^{-3/2} U_{\mathcal{I}}(n, \alpha) \xrightarrow[n \rightarrow \infty]{\mathbb{R}} U_{\mathcal{I},1}(\alpha). \quad (11)$$

Here

$$U_{\mathcal{I},1}(\alpha) = \sup_{t,s} \frac{|J(t) - J(s) - (t-s)J(1)|}{|(t-s)(1-(t-s))|^\alpha}$$

where $J(t) = \int_0^t U_\gamma(r) dr$ and $U_\gamma(r) = \int_0^s e^{\gamma(r-v)} dW(v)$.

Consistency of the test in case 1

Theorem

Let $0 < \alpha < 1/2$. Suppose (y_k) is generated by (4) and $\phi_n = e^{\gamma/n}$, where $\gamma < 0$ is a constant. Under alternative H_A , if

$$\lim_{n \rightarrow \infty} |a_n| h_n^{1-\alpha} n^{-1/2} = \infty, \quad \text{where} \quad h_n := \frac{l^*}{n} \left(1 - \frac{l^*}{n} \right)$$

then

$$n^{-3/2} U_{IS}(n, \alpha) \xrightarrow[n \rightarrow \infty]{P} \infty.$$

Behaviour under null hypothesis in case 2

Theorem

Suppose (y_k) is generated by (4) and $\phi_n = 1 - \gamma_n/n$, where (γ_n) is a sequence of non negative numbers, $\gamma_n \rightarrow \infty$ and $\gamma_n/n \rightarrow 0$, as $n \rightarrow \infty$. Assume also that the innovations (ε_k) are i.i.d. and satisfy condition $\lim_{t \rightarrow \infty} t^p P(|\varepsilon_0| > t) = 0$ for some $p > 2$. Put $\alpha = \frac{1}{2} - \frac{1}{p}$. Also for $0 < \beta < \alpha$ condition $\liminf_{n \rightarrow \infty} \gamma_n n^{-\frac{\beta}{\alpha}} > 0$ holds. Then under hypothesis H_0

$$n^{-1/2}(1 - \phi_n)U_{1S}(n, \alpha) \xrightarrow[n \rightarrow \infty]{\mathbb{R}} U_{12}(\alpha),$$

Here

$$U_{12}(\alpha) = \sup_{t,s} \frac{|B(t) - B(s)|}{|(t-s)(1-(t-s))|^\alpha}$$

where $B(t) = W(t) - tW(1)$ is Brownian bridge.

Consistency of the test in case 2

Theorem

Let $0 < \alpha < 1/2$. Suppose (y_k) is generated by (4) and $\phi_n = 1 - \gamma_n/n$, where (γ_n) is a sequence of non negative numbers, $\gamma_n \rightarrow \infty$ and $\gamma_n/n \rightarrow 0$, as $n \rightarrow \infty$. Under alternative H_A , if

$$\lim_{n \rightarrow \infty} |a_n| h_n^{1-\alpha} n^{-1/2} \gamma_n = \infty, \quad \text{where} \quad h_n := \frac{l^*}{n} \left(1 - \frac{l^*}{n}\right)$$

then

$$(1 - \phi_n) n^{-1/2} UI_S(n, \alpha) \xrightarrow[n \rightarrow \infty]{P} \infty.$$

Test statistics with $\hat{\epsilon}_k$'s

We construct uniform increment statistics :

$$Ul_W(n, \alpha) := \sup_{k^*, l^*} \frac{\left| \widehat{W}_n^{\text{st}}(\mathbb{I}_n^*) - \frac{l^*}{n} \widehat{W}_n^{\text{st}}(n) \right|}{\left| \frac{l^*}{n} \left(1 - \frac{l^*}{n} \right) \right|^\alpha}$$

where

$$\begin{aligned} \mathbb{I}_n^* &= \{k^* + 1, \dots, k^* + l^*\} \\ \widehat{W}_n^{\text{st}}(\mathbb{I}_n^*) &= \sum_{k \in \mathbb{I}_n^*} \hat{\epsilon}_k \\ \widehat{W}_n^{\text{st}}(n) &= \sum_{k=1}^n \hat{\epsilon}_k \end{aligned}$$

Behaviour under null hypothesis in case 1

Theorem

Suppose that (y_k) is generated by (4), $\phi_n = e^{\gamma/n}$ with $\gamma < 0$ and that the sequence (ε_k) are i.i.d. random variables with mean 0 and variance 1. Coefficient $\hat{\phi}_n$ is estimated by LSE. Then under hypothesis H_0

$$n^{-1/2} U_{W,n}(\alpha) \xrightarrow[n \rightarrow \infty]{\mathbb{R}} U_{W,1}(\alpha). \quad (12)$$

Here

$$U_{W,1}(\alpha) = \sup_{t,s} \frac{|B(t) - B(s) - A^{-1}B'(J(t) - J(s) - (t-s)J(1))|}{|(t-s)(1-(t-s))|^\alpha}$$

Here $B' = \int_0^1 U_\gamma(r) dW(r)$, $A = \int_0^1 U_\gamma(r)^2 dr$ and $J(t) := \int_0^t U_\gamma(s) ds$, $B(t)$ is a Brownian bridge.

Consistency of the test in case 1

Theorem

Let $0 < \alpha < 1/2$. Suppose (y_k) is generated by (4) and $\phi_n = e^{\gamma/n}$, where $\gamma < 0$ is a constant. Under alternative H_A , if

$$\lim_{n \rightarrow \infty} |a| h_n^{1-\alpha} n^{1/2} = \infty, \quad \text{where } h_n := \frac{l^*}{n} \left(1 - \frac{l^*}{n}\right)$$

then

$$n^{-1/2} Ul_W(n, \alpha) \xrightarrow[n \rightarrow \infty]{P} \infty.$$

Behaviour under null hypothesis in case 2

Theorem

Suppose (y_k) is generated by (4) and $\phi_n = 1 - \gamma_n/n$, where (γ_n) is a sequence of non negative numbers, $\gamma_n \rightarrow \infty$ and $\gamma_n/n \rightarrow 0$, as $n \rightarrow \infty$. Assume also that the innovations (ε_k) are i.i.d. and satisfy condition $\lim_{t \rightarrow \infty} t^p P(|\varepsilon_0| > t) = 0$ for some $p > 2$. Put $\alpha = \frac{1}{2} - \frac{1}{p}$. Also for $0 < \beta \leq \alpha$ condition $\liminf_{n \rightarrow \infty} \gamma_n n^{-\frac{2\beta}{1+2\alpha}} > 0$ holds. Then under hypothesis H_0

$$n^{-1/2} U_{IW}(n, \alpha) \xrightarrow[n \rightarrow \infty]{\mathbb{R}} U_2(\alpha),$$

Here

$$U_2(\alpha) = \sup_{t,s} \frac{|B(t) - B(s)|}{|(t-s)(1-(t-s))|^\alpha}$$

where $B(t) = W(t) - tW(1)$ is Brownian bridge.

Consistency of the test in case 2

Theorem

Let $0 < \alpha < 1/2$. Suppose (y_k) is generated by (4) and $\phi_n = 1 - \gamma_n/n$, where (γ_n) is a sequence of non negative numbers, $\gamma_n \rightarrow \infty$ and $\gamma_n/n \rightarrow 0$, as $n \rightarrow \infty$. Under alternative H_A , if

$$\lim_{n \rightarrow \infty} |a| h_n^{1-\alpha} n^{1/2} = \infty, \quad \text{where} \quad h_n := \frac{l^*}{n} \left(1 - \frac{l^*}{n}\right)$$

then

$$n^{-1/2} Ul_W(n, \alpha) \xrightarrow[n \rightarrow \infty]{P} \infty.$$

AČIŪ UŽ DĒMESĪ :))